

# 11. CONDITIONAL EXPECTATION

## DEFINITION.

Given  $(\Omega, \mathcal{F}, \mathcal{P})$  and a  $\sigma$ -algebra  $\mathcal{G} \subset \mathcal{F}$ , a random variable  $X(\omega)$  on  $\Omega$ , we define the **CONDITIONAL**

**EXPECTATION OF  $X$  GIVEN  $\mathcal{G}$**

the unique (a.e.) random variable

$Y(\omega) = E\{X | \mathcal{G}\}(\omega)$  such that

- it is  $\mathcal{G}$ -measurable, i.e.

$$Y^{-1}(B) \in \mathcal{G} \quad \forall B \in \mathcal{B}(\mathbb{R})$$

- $\int_A X(\omega) d\mathcal{P}(\omega) = \int_A E\{X | \mathcal{G}\}(\omega) d\mathcal{P}(\omega)$

$$\forall A \in \mathcal{G}.$$

EXAMPLE. Given a random variable  $X(\omega)$  on  $(\Omega, \mathcal{F}, \mathcal{P})$ , calculate its conditional expectation given  $\mathcal{G} = \{\emptyset, A, A^c, \Omega\}$  where  $A \in \mathcal{F}$ . The events  $\{\emptyset, A, A^c, \Omega\}$  are the atoms of  $\mathcal{G}$ . Since any random variable which is  $\mathcal{G}$ -measurable must be constant over the atoms and  $E\{X | \mathcal{G}\}$  must be  $\mathcal{G}$ -measurable,  $E\{X | \mathcal{G}\}$  must have the form

$$E\{X | \mathcal{G}\}(\omega) = \alpha_1 \chi_A(\omega) + \alpha_2 \chi_{A^c}(\omega)$$

By the definition of  $E\{X | \mathcal{G}\}$

$$\int_A E\{X | \mathcal{G}\} d\mathcal{P} = \alpha_1 \mathcal{P}(A) = \int_A X d\mathcal{P}$$

$$\int_{A^c} E\{X | \mathcal{G}\} d\mathcal{P} = \alpha_2 \mathcal{P}(A^c) = \int_{A^c} X d\mathcal{P}$$

It follows

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$$E\{X|G\}(\omega) = \begin{cases} \frac{1}{P(A)} \int_A X(\omega) dP(\omega) & \text{if } \omega \in A \\ \frac{1}{P(A^c)} \int_{A^c} X(\omega) dP(\omega) & \text{if } \omega \in A^c \end{cases}$$

In general, if  $G$  is generated by a countable family of atoms  $\{A_i\}$ , it is possible to see

$$E\{X|A\}(\omega) = \sum_{i=1}^{\infty} \left( \frac{1}{P(A_i)} \int_{A_i} X(\omega) dP(\omega) \right) \chi_{A_i}(\omega)$$

Notice also that

$$\begin{aligned} \int_{A_i} X(\omega) dP(\omega) &= \int_{\Omega} \chi_{A_i}(\omega) X(\omega) dP(\omega) \\ &= E\{ \chi_{A_i} X(\omega) \} \end{aligned} \quad \text{in so that}$$

$$E\{X|A\}(\omega) = \sum_{i=1}^{\infty} \frac{1}{P(A_i)} E\{ \chi_{A_i} X(\omega) \} \chi_{A_i}(\omega) \quad \blacktriangleleft$$

EXAMPLE Consider  $(\Omega, \mathcal{F}, \mathbb{P})$

with  $\Omega = [0, 1]$ ,  $\mathbb{P} : \mathbb{P}([a, b]) = b - a$  and  $\mathcal{F}$  is the  $\sigma$ -algebra generated by the following atoms

$$\{ \emptyset, [0, 1/3), [1/3, 1/2), [1/2, 4/5), [4/5, 1] \}$$

Any random variable  $X(\omega)$  on  $(\Omega, \mathcal{F}, \mathbb{P})$  is constant on the atoms. Consider

for example

$$X(\omega) = \begin{cases} 5 & \omega \in [0, 1/3) \\ 2 & \omega \in [1/3, 1/2) \\ 3 & \omega \in [1/2, 4/5) \\ 0.5 & \omega \in [4/5, 1] \end{cases}$$

and the  $\sigma$ -algebra  $\mathcal{G} \subset \mathcal{F}$

generata dagli atomi

$$\mathcal{G} = \{ \emptyset, [0, 1/2), [1/2, 1] \}$$

Let calculate  $E\{X|G\}$  on the atom  $[0, 1/2)$  :

$$E\{X|G\}(\omega) = \frac{1}{P([0, 1/2))} \int_{[0, 1/2)} X(\omega) dP(\omega)$$

$$= \frac{1}{1/2} \cdot [5 \cdot P([0, 1/3)) + 2 \cdot P([1/3, 1/2))]$$

$$= 4 \quad \text{if } \omega \in [0, 1/2)$$

and

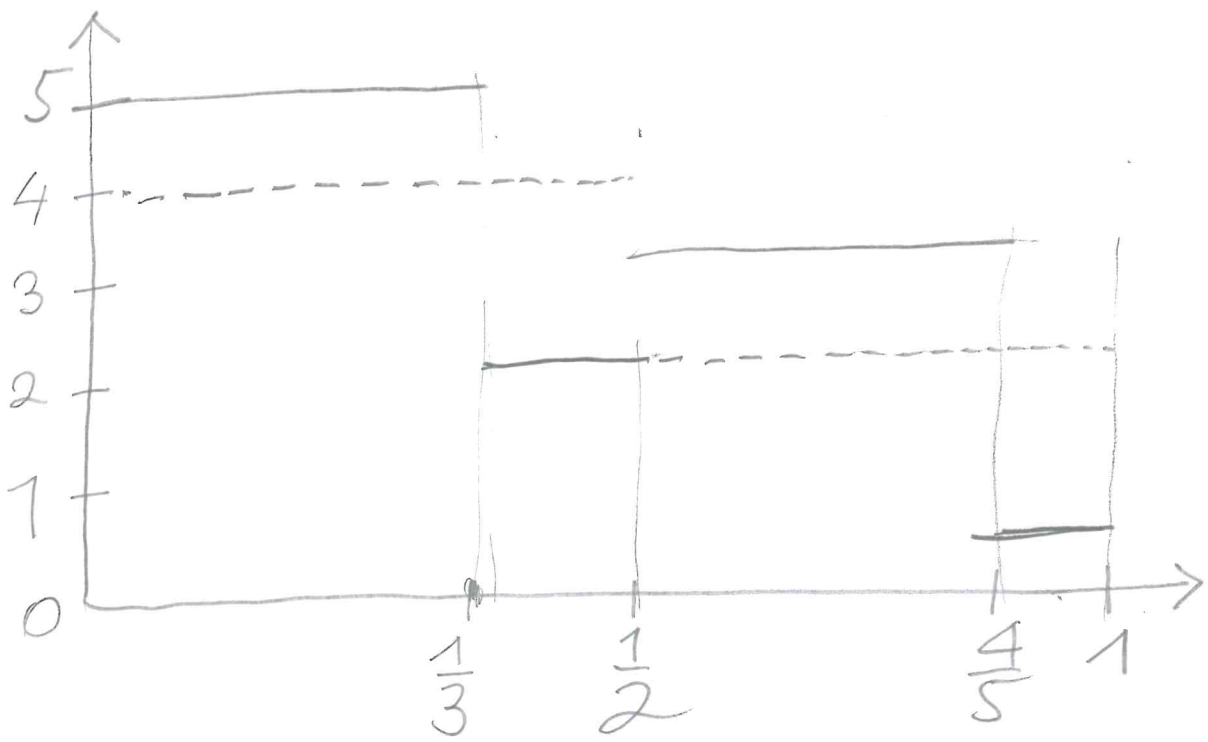
$$E\{X|G\}(\omega) = \frac{1}{P([1/2, 1])} \int_{[1/2, 1]} X(\omega) dP(\omega)$$

$$= \frac{1}{1/2} \cdot [3 \cdot P([1/2, 4/5)) + 0.5 P([4/5, 1])]$$

$$= 2 \quad \text{if } \omega \in [1/2, 1]$$

Notice that

$$E\{X\} = \int X(\omega) dP(\omega) = 5 \cdot P([0, 1/3)) + 2 \cdot P([1/3, 1/2)) + 3 \cdot P([1/2, 4/5)) + \frac{1}{2} P([4/5, 1]) = 3$$



$X(\omega)$  ———  
 $E\{X|G\}(\omega)$  - - -

$E\{X|G\}$  is the (weighted) mean of  $X$  over  $[0, 1/2)$  and  $[1/2, 1]$  respectively!

DEFINITION. For given random variables  $X$  and  $\{Y_1, \dots, Y_n\}$  over  $(\Omega, \mathcal{F}, \mathcal{P})$  we define the CONDITIONAL EXPECTATION OF  $X$  given  $\{Y_1, \dots, Y_n\}$  the conditional expectation of  $X$  given the  $\sigma$ -algebra  $\mathcal{G}$  of  $Y_1, \dots, Y_n$  generated by  $\{Y_1, \dots, Y_n\}$ . We write  $E\{X | Y_1, \dots, Y_n\}$ .

EXAMPLE.  $\Omega = [0, 1]$ ,  $\mathcal{F}$  is the Borel  $\sigma$ -algebra on  $\Omega = [0, 1]$ ,  $\mathcal{P} : \mathcal{P}([a, b]) = b - a$ . Consider

$$X(\omega) = \omega$$

and

$$Y_1(\omega) = \begin{cases} 1 & \omega \in [0, 1/2) \\ 2 & \omega \in [1/2, 1] \end{cases} \quad Y_2(\omega) = \begin{cases} 3 & \omega \in [0, 1/3] \\ 2 & \omega \in (1/3, 2/3) \\ 4 & \omega \in [2/3, 1] \end{cases}$$

Clearly,  $\mathcal{G}_{Y_1, Y_2} \subset \mathcal{F}$  and  $\mathcal{G}_{Y_1, Y_2}$  is the smallest  $\sigma$ -algebra containing

$$\left\{ [0, 1/3], (1/3, 1/2), [1/2, 2/3), [2/3, 1] \right\}.$$



Since  $E\{X | Y_1, Y_2\}$  is  $\mathcal{F}_{Y_1, Y_2}$ -measurable, it is constant on the atoms  $\{[0, 1/3], (1/3, 1/2), [1/2, 2/3), [2/3, 1]\}$ .

We have

$$E\{X | Y_1, Y_2\} = \begin{cases} 1/6 & \omega \in [0, 1/3) \\ 5/12 & \omega \in (1/3, 1/2) \\ 7/12 & \omega \in [1/2, 2/3) \\ 5/6 & \omega \in [2/3, 1] \end{cases}$$

### 11.1 PROPERTIES OF CONDITIONAL EXPECTATION

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i) if  $X$  is random variable which is also  $\mathcal{G}$ -measurable:

$$E\{X | \mathcal{G}\} = X \text{ a.e.}$$

ii) if  $\mathcal{F}_m = \{\emptyset, \Omega\}$

$$E\{X | \mathcal{G}\} = E\{X\}$$



iii)  $\forall a_1, a_2 \in \mathbb{R}$ ,

$X_1, X_2$  random variables :

$$E\{a_1 X_1 + a_2 X_2 \mid \mathcal{G}\} = a_1 E\{X_1 \mid \mathcal{G}\} + a_2 E\{X_2 \mid \mathcal{G}\}$$

iv) if  $X_1, X_2$  are random variables  
and  $X_1 \leq X_2$  a.e. :

$$E\{X_1 \mid \mathcal{G}\} \leq E\{X_2 \mid \mathcal{G}\} \text{ a.e.}$$

v) if  $X$  is a random variable

$$E\{E\{X \mid \mathcal{G}\}\} = E\{X\}.$$

REMARK Hölder, Cauchy, Minkowski,  
Jensen inequalities straightforwardly  
extend to the case of conditional  
expectations  $\blacktriangleleft$

FACT. Given random variables  $X, Y$  on  $(\Omega, \mathcal{F}, \mathbb{P})$  and  $Y$  is  $\mathcal{G}$ -measurable,  $\mathcal{G} \subset \mathcal{F}$ ,

$$E\{XY | \mathcal{G}\} = Y E\{X | \mathcal{G}\} \text{ a.e.}$$

FACT. Given  $\sigma$ -algebras  $\mathcal{G}_1, \mathcal{G}_2$  with  $\mathcal{G}_1 \subset \mathcal{G}_2 \subset \mathcal{F}$  and  $X$  random variable on  $(\Omega, \mathcal{F}, \mathbb{P})$ :

$$\begin{aligned} E\{X | \mathcal{G}_1\} &= E\{E\{X | \mathcal{G}_1\} | \mathcal{G}_2\} \\ &= E\{E\{X | \mathcal{G}_2\} | \mathcal{G}_1\} \text{ a.e.} \end{aligned}$$

FACT. There exists a measurable function  $f: \mathbb{R} \rightarrow \mathbb{R}$  such

$$E\{X | Y\}(\omega) = f(Y(\omega))$$

The above fact proves that from the values of  $Y(\omega)$  it is possible to obtain the expectation of  $X(\omega)$ .

$$\begin{aligned}
&= 1 \cdot P(A \cap \{\omega_{13}, \omega_{5}\}) / P(A) \\
&+ 3 \cdot P(A \cap \{\omega_2, \omega_4\}) / P(A) \\
&+ 2 \cdot P(A \cap \{\omega_3\}) / P(A) = \frac{6}{5}
\end{aligned}$$

Finally, notice that  $X$  and  $Y$  are not independent:

$$P\{X=1\} = P\{\{\omega_{13}, \omega_5\}\} = \frac{2}{16}$$

$$P\{Y=1\} = P\{\{\omega_1, \omega_3\}\} = \frac{2}{8}$$

$$P\{X=1, Y=1\} = P\{\{\omega_1\}\} = \frac{1}{16}$$

But  $P\{X=1\}P\{Y=1\} = \frac{4}{128}$

$$\neq P\{X=1, Y=1\} = \frac{1}{16} \quad \blacktriangleleft$$

As already seen,

$$\begin{aligned}
 & X, Y \text{ independent} \Rightarrow \\
 & E\{X|Y\} = E\{X\}
 \end{aligned}$$

The converse is false but

$$\begin{aligned}
 & E\{X|Y\} = E\{X\} \Rightarrow \\
 & X, Y \text{ uncorrelated}
 \end{aligned}$$

Indeed,

$$\begin{aligned}
 \sigma_{XY} &= E\{(Y - E\{Y\})(X - E\{X\})\} = \\
 &= E\{YX\} - E\{Y\}E\{X\} = \\
 &= E\{E\{YX|Y^Y\}\} - E\{Y\}E\{X\} = \\
 &= E\{YE\{X|Y^Y\}\} - E\{YE\{X\}\} = \\
 &= E\{Y(E\{X|Y^Y\} - E\{X\})\} = 0
 \end{aligned}$$

It is also possible to write  $E\{X|Y\}$  as some function of  $Y$ :

$$E\{X|Y\}(y) = \begin{cases} 2 & y=0 \\ 6/5 & y=1 \\ 3 & y=2 \end{cases}$$

If we take  $\mathcal{P}\{\cdot|Y\}$  as a probability measure on  $(\Omega, \mathcal{F})$  we can calculate the conditional expectation  $E\{X|Y\}$  exactly as we did for  $E\{X\}$  with probability measure  $\mathcal{P}$ :

$$E\{X|Y\} = \int_{\Omega} X(\omega) d\mathcal{P}\{\omega|Y\}$$

For example, for  $\omega \in \mathcal{A}$ :

$$\begin{aligned} E\{X|Y\}(\omega) &= 1 \cdot \mathcal{P}\{\{\omega_1, \omega_5\}|Y\}(\omega) \\ &+ 3 \cdot \mathcal{P}\{\{\omega_2, \omega_4\}|Y\}(\omega) \\ &+ 2 \cdot \mathcal{P}\{\{\omega_3\}|Y\}(\omega) = \end{aligned}$$

Therefore, for calculating  $E\{X|Y\}$  it is sufficient to calculate it on the atoms of  $\mathcal{F}$ . 79

We obtain

$$\omega \in A: \int_A X dP = \frac{3}{4} \Rightarrow E\{X|Y\} = \frac{6}{5}$$

$$\omega \in B: \int_B X dP = \frac{1}{4} \Rightarrow E\{X|Y\} = 3$$

$$\omega \in C: \int_C X dP = \frac{1}{4} \Rightarrow E\{X|Y\} = 2$$

Let's see how to determine  $P(D|Y)$  for any  $D \in \mathcal{F}$ . Also in this case, it is sufficient to calculate  $P(D|Y)$  on the atoms of  $\mathcal{F}$ :

$$\omega \in A: \int_A X dP = P(A \cap D) \Rightarrow P(D|Y) = \frac{P(A \cap D)}{P(A)}$$

$$\omega \in B: \int_B X dP = P(B \cap D) \Rightarrow P(D|Y) = \frac{P(B \cap D)}{P(B)}$$

$$\omega \in C: \int_C X dP = P(C \cap D) \Rightarrow P(D|Y) = \frac{P(C \cap D)}{P(C)}$$



EXAMPLE.  $\Omega = \{\omega_1, \dots, \omega_5\}$

with  $\mathcal{F} = \mathcal{F}_M$  and

$$P\{\omega_1\} = \frac{1}{2}, P\{\omega_2\} = \frac{1}{4}, P\{\omega_3\} = \frac{1}{8},$$

$$P\{\omega_4\} = \frac{1}{16}, P\{\omega_5\} = \frac{1}{16}.$$

Let

$$X(\omega) = \begin{cases} 1 & \omega \in \{\omega_1, \omega_5\} \\ 3 & \omega \in \{\omega_2, \omega_4\} \\ 2 & \omega = \omega_3 \end{cases}$$

$$Y(\omega) = \begin{cases} 1 & \omega \in \{\omega_1, \omega_3\} \\ 2 & \omega = \omega_2 \\ 0 & \omega \in \{\omega_4, \omega_5\} \end{cases}$$

Notice that

$$\mathcal{F}^Y = \{\emptyset, A, B, C, A^c, B^c, C^c, \Omega, A \cup B, B \cup C, A \cup C\}$$

with  $A = \{\omega_1, \omega_3\}$ ,  $B = \omega_2$ ,  $C = \{\omega_4, \omega_5\}$ .

The atoms of  $\mathcal{F}^Y$  are  $A, B, C$ .

We calculate  $E\{X|Y\}$ .

For  $\omega \in B = \{\omega_1, \omega_3\}$ :

$$E\{X|Y\} = \frac{1}{P(B)} \int_B X dP = \frac{7}{3}$$

and  $\omega \in B^c = \{\omega_2, \omega_4\}$ :

$$E\{X|Y\} = \frac{1}{P(B^c)} \int_{B^c} X dP = \frac{7}{3}$$

Therefore,  $E\{X|Y\} = \frac{7}{3} \forall \omega$ . Moreover,

$$E\{X\} =$$

$$= P\{\omega_1, \omega_2\} \cdot 1 + P\{\omega_3, \omega_4\} \cdot 3$$

$$= \frac{1}{3} \cdot 1 + \frac{2}{3} \cdot 3 = \frac{7}{3} = E\{X|Y\}$$

It is straightforward to see that  $X$  and  $Y$  are independent since  $\mathcal{F}^X$  and  $\mathcal{F}^Y$  are  $\blacktriangleleft$

Therefore, in this sense the conditional expectation  $E\{X|Y\}$  gives evidence of the correlation between  $X$  and  $Y$ . A fact which further support this interpretation is :

FACT. If  $X$  and  $Y$  are independent random variables :

$$E\{X|Y\}(\omega) = X(\omega) \text{ a.e.}$$

It is possible also to define a **CONDITIONAL PROBABILITY** measure and a **CONDITIONAL PROBABILITY DENSITY**.

DEFINITION. Given  $(\Omega, \mathcal{F}, \mathbb{P})$  and a  $\sigma$ -algebra  $\mathcal{G} \subset \mathcal{F}$ , we define **CONDITIONAL PROBABILITY OF  $A \in \mathcal{F}$ , GIVEN  $\mathcal{G}$** , as the random variable

$$P(A|\mathcal{G}) \triangleq E\{X_A|\mathcal{G}\}.$$

Similarly, we define the **CONDITIONAL PROBABILITY** OF  $A \in \mathcal{F}$ , given the random variable  $Y$ , as

$$\begin{aligned}
 P(A|Y) &= P(A|\mathcal{F}^Y) \\
 &\triangleq E\{X_A | \mathcal{F}^Y\}
 \end{aligned}$$

EXAMPLE. Consider  $\Omega = \{\omega_1, \omega_2, \omega_3, \omega_4\}$

and  $\mathcal{F} = \mathcal{F}_M = \{ \text{all subset of } \Omega \}$

with

$$P\{\omega_1\} = \frac{1}{9}, \quad P\{\omega_2\} = \frac{2}{9}, \quad P\{\omega_3\} = \frac{2}{9}, \quad P\{\omega_4\} = \frac{4}{9}$$

Moreover,

$$X(\omega) = \begin{cases} 1 & \omega \in \{\omega_1, \omega_2\} \\ 3 & \omega \in \{\omega_3, \omega_4\} \end{cases}$$

$$Y(\omega) = \begin{cases} 1 & \omega \in \{\omega_1, \omega_3\} \\ 2 & \omega \in \{\omega_2, \omega_4\} \end{cases}$$

As usual,

$$\mathcal{F}^X = \{\emptyset, A, A^c, \Omega\}, \quad \mathcal{F}^Y = \{\emptyset, B, B^c, \Omega\}$$

where  $A = \{\omega_1, \omega_2\}$  and  $B = \{\omega_1, \omega_3\}$ .

## 12. CONDITIONAL PROBABILITY DENSITY AND BAYES FORMULAS

We have introduced in the previous section the notion of conditional expectation of  $X$  given  $Y$ : this notion well-characterizes to what extent the a posteriori knowledge of the results of  $Y$ , improves our a priori knowledge of the results of  $X$ . As we did for expectation of  $X$ , we want to characterize a conditional probability density: to evaluate conditional expectations.

EXAMPLE  $\Omega = \{\omega_1, \dots, \omega_n\}$ ,

$$P(\omega_i) = \frac{1}{n} \quad \forall \omega_i \in \Omega, \quad \mathcal{F} = \mathcal{F}_M.$$

Consider  $A, B \in \mathcal{F}$  with cardinality  $\alpha$  and  $\beta$ , respectively, such that  $A \cap B$  contains  $\gamma$  points  $\omega_i$ . Therefore,

$$P(A) = \frac{\alpha}{n}, \quad P(B) = \frac{\beta}{n}, \quad P(A \cap B) = \frac{\gamma}{n}$$

If we know the result of our experiment is in  $B$ , what is the probability that the result is also in  $A$ ? Clearly, we may naturally consider  $B$  (instead of  $\Omega$ ) as the new space of results with some probability

$$P(\omega_i | B) = \frac{1}{\beta}$$

for each  $\omega_i \in B$ . We have on  $A$ :



$$P(A|B) = \frac{\alpha}{\beta}$$

$$= \frac{\alpha/n}{\beta/n} = \frac{P(A \cap B)}{P(B)} \quad \triangleleft$$

DEFINITION. We define CONDITIONAL PROBABILITY of  $A$  given  $B$ , denoted by  $P(A|B)$  :

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

with  $P(B) \neq 0$ .

If  $A$  and  $B$  are independent :

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A)P(B)}{P(B)} = P(A)$$

i.e. the a posteriori knowledge that the event  $B$  happened does not improve our a priori knowledge on the possible happening of  $A$ .

From our definition of  $P(A|B)$  it is possible to define a conditional expectation of a random variable  $X$  on  $(\Omega, \mathcal{F}, P)$  given that  $B \in \mathcal{F}$  happened:

$$E\{X(\omega)|B\} \triangleq \frac{E\{X(\omega)X_B(\omega)\}}{P(B)},$$

$$P(B) \neq 0.$$

Notice if  $X = X_A$

$$\begin{aligned} E\{X_A|B\} &= \frac{E\{X_A X_B\}}{P(B)} = \frac{P(A \cap B)}{P(B)} \\ &= P(A|B) \end{aligned}$$

Straightforwardly, we can extend our definition to the expectation of  $X$  given the values of another random variable  $Y$  in  $B \in \mathcal{F}$ :

$$E\{X(\omega) | Y \in B\} = \frac{E\{X \chi_{Y^{-1}(B)}(\omega)\}}{P(Y^{-1}(B))}$$

$$= \int \frac{X(\omega) \chi_{Y^{-1}(B)}(\omega)}{P(Y^{-1}(B))} dP(\omega)$$

$$= \frac{\int_{Y^{-1}(B)} X(\omega) dP(\omega)}{P(Y^{-1}(B))} = \int_{Y^{-1}(B)} E\{X(\omega) | Y\}(\omega) dP(\omega | B)$$

where  $P(\omega | B) = \frac{P(\omega \cap Y^{-1}(B))}{P(Y^{-1}(B))}$

if  $P(Y^{-1}(B)) \neq 0$ . This coincides with our previous definition of  $E\{X(\omega) | Y\}$  (resp.  $P\{A | Y\}$ ) (see previous section) in the sense that:

$$E\{X(\omega) | Y \in B\} = E\{X(\omega) | Y\}(\omega)$$

$$P\{A | B\} = P\{A | Y\}(\omega)$$

$$\forall \omega \in Y^{-1}(B), A \in \mathcal{F}.$$

Notice that the above definitions depend on the fact that  $P(B) \neq 0$  or  $P(Y^{-1}(B)) \neq 0$ . To circumvent this problem and obtain a general definition of conditional probability density we resort to a fundamental result of Bayes: given two random variables  $X$  and  $Y$  exists a unique (a.e.) function  $h(x, y)$  such that

$$E\{X(\omega) \mid Y\} \stackrel{F}{=} \int_{Y(\omega)=y} x h(x, y) dx$$

It is quite natural to take  $h(x, y)$  as conditional probability density of  $X$  given  $Y$ .

From this it is possible to prove: Pg

BAYES THEOREM. Given random variables  $X, Y$  and its joint probability density  $f_{X,Y}(x,y)$ , the conditional probability density  $f_{X|Y}(x,y)$  of  $X$  given  $Y$  and, respectively,  $f_{Y|X}(y,x)$  of  $Y$  given  $X$  satisfy:

$$f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$f_{Y|X}(y,x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

Notice that

$$f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{Y|X}(y,x) f_X(x)}{\int_{\mathbb{R}} f_{Y|X}(y,x) f_X(x) dx}$$

using the formulas for marginal densities.

Conditional density can be used to evaluate conditional expectations given a posteriori values of the conditioning random variable:

$$E\{X|Y\}_{Y=y} = \int_{\mathbb{R}} x p_{X|Y}(x,y) dx$$

$$E\{Y|X\}_{X=x} = \int_{\mathbb{R}} y p_{Y|X}(y,x) dy$$



## 12. CONDITIONING AND ORTHOGONAL PROJECTIONS

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Given  $Y \in L_2(\Omega, \mathcal{F}, \mathbb{P})$ ,  
 consider the set of  $\mathcal{F}^Y$ -measurable  
 random variables, i.e. the  
 set of functions  $Z = f(Y)$  for  
 some  $\mathcal{B}(\mathbb{R})$ -measurable function  $f: \mathbb{R} \rightarrow \mathbb{R}$ . This is a linear space  
 (which we denote by  $\mathcal{M}^Y$ ).

We want to see that the ortho-  
 gonal projection of a random va-  
 riable  $X$  on  $\mathcal{M}^Y$  (denoted  $\Pi(X | \mathcal{M}^Y)$ )  
 is exactly  $E\{X | \mathcal{F}^Y\}$ . Since

$$X = X_{\parallel} + X_{\perp}$$

where  $X_{\parallel}, X_{\perp}$  are the unique  
 random variables such that  $X_{\parallel} \in \mathcal{M}^Y$   
 and  $X_{\perp} \in (\mathcal{M}^Y)^{\perp}$  (the orthogonal of  $\mathcal{M}^Y$ ),

it is sufficient to prove that

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$$i) X - E\{X | \mathcal{F}^Y\} \perp \in (\mathcal{M}^Y)^\perp$$

$$ii) E\{X | \mathcal{F}^Y\} \in \mathcal{M}^Y.$$

ii) is true because  $E\{X | \mathcal{F}^Y\}$  is  $\mathcal{F}^Y$ -measurable.

We prove i). To prove this, we

prove

$$\langle X - E\{X | \mathcal{F}^Y\}, Z \rangle = 0 \\ \forall Z \in \mathcal{M}^Y.$$

Indeed if  $Z \in \mathcal{M}^Y$ :

$$\begin{aligned} \langle X - E\{X | \mathcal{F}^Y\}, Z \rangle &= E\{(X - E\{X | \mathcal{F}^Y\})Z\} \\ &= E\{XZ - E\{XZ | \mathcal{F}^Y\}\} = \\ &= E\{XZ\} - E\{E\{XZ | \mathcal{F}^Y\}\} = \\ &= E\{XZ\} - E\{XZ\} = 0 \end{aligned}$$

The orthogonal projection has the following important property. 93

### PROJECTION THEOREM. Let

$H$  a linear space with a scalar product  $\langle \cdot, \cdot \rangle_H$ ,  $\mathcal{M} \subset H$  a closed subspace of  $H$ .  $\forall v \in H \exists!$

$m_0 \in \mathcal{M}$  :

$$(*) \quad \|v - m_0\|_H \leq \|v - m\|_H, \quad \forall m \in \mathcal{M},$$

where  $\|\cdot\|_H = \langle \cdot, \cdot \rangle$ . Necessary and sufficient condition for  $m_0$  being the unique element of  $\mathcal{M}$  satisfying (\*) is

$$\langle v - m_0, m \rangle_H = 0 \quad \forall m \in \mathcal{M}.$$

In other words,  $\forall v \in H$

$$\arg \min_{m \in \mathcal{M}} \|v - m\| = \Pi(v | \mathcal{M})$$